

GENERALIZED RESTRICTED ISOMETRY PROPERTY FOR α -STABLE RANDOM PROJECTIONS

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- Motivation
- α -stable random projections
- Restricted- p Isometry Property (RIP_p)
- Computer simulations
- Conclusions

In order to show that a sparse signal can be recovered exactly from a limited number of measurements, Candès and Tao introduced the restricted isometry property (RIP)[†]:

Let \mathbf{R} be an $K \times N$ matrix, where $K < N$, and s a positive integer. Then δ_s is the smallest number such that

$$(1 - \delta_s) \|x\|_2^2 \leq \|\mathbf{R}x\|_2^2 \leq (1 + \delta_s) \|x\|_2^2$$

for all x such that $\|x\|_0 \leq s$.

Here, δ_s is a number that measures how much \mathbf{R} acts like an isometry on s -sparse vectors.

[†] Candès E. and Tao T., "Decoding by linear programming", IEEE Trans. Inf. Theory 51, 4203-4215, 2005.

Chartrand and Staneva proposed a variant of the RIP to show that l_p minimization may be used to recover sparse signals[†]:

Let \mathbf{R} be an $K \times N$ matrix, where $K < N$, L a positive integer and $0 < p \leq 1$. Then δ_L is the smallest number such that

$$(1 - \delta_L) \|x\|_2^p \leq \|\mathbf{R}x\|_2^p \leq (1 + \delta_L) \|x\|_2^p$$

for all x such that $\|x\|_0 \leq L$.

They used this result to show that l_p minimization requires fewer measurements than l_1 minimization. In this case, the entries of \mathbf{R} are Gaussian realizations.

[†] R. Chartrand and V. Staneva, "Restricted isometry properties and nonconvex compressed sensing", Inverse Problems, vol. 24, no. 3, June 2008.

The latter result can be extended to \mathbf{R} having α -stable distributions if $1 \leq \alpha \leq 2$:

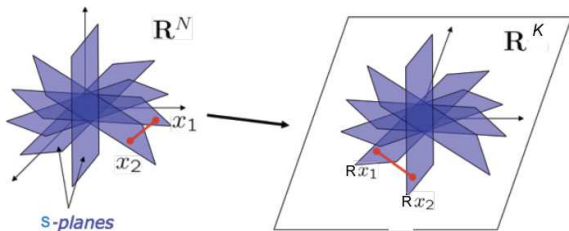
Let \mathbf{R} be an $K \times N$ matrix, where $K < N$, s a positive integer and $p < \alpha/2$. Then, for a given δ

$$(1 - \delta) \|x\|_{\alpha}^p < \|\mathbf{R}x\|_p^p < (1 + \delta) \|x\|_{\alpha}^p$$

for all x such that $\|x\|_0 = s$.

Despite that the variance of α -stable variables is infinite, this RIP may hold with probability greater than $1 - 1/\binom{N}{s}$.

MOTIVATION



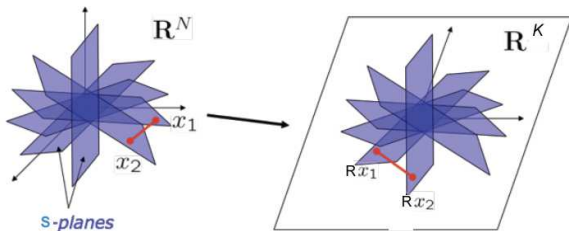
ℓ_2 based dimensionality reduction:

$$\|\mathbf{R}x_2 - \mathbf{R}x_1\|_{\ell_2}^2 \approx \|x_2 - x_1\|_{\ell_2}^2$$

ℓ_1 based dimensionality reduction:

$$\|\mathbf{R}x_2 - \mathbf{R}x_1\|_{\ell_p}^p \approx \|x_2 - x_1\|_{\ell_1}^p; p < 1/2$$

MOTIVATION



ℓ_2 based dimensionality reduction:

$$\|\mathbf{R}x_2 - \mathbf{R}x_1\|_{\ell_2}^2 \approx \|x_2 - x_1\|_{\ell_2}^2$$

ℓ_1 based dimensionality reduction:

$$\|\mathbf{R}x_2 - \mathbf{R}x_1\|_{\ell_p}^p \approx \|x_2 - x_1\|_{\ell_1}^p; \quad p < 1/2$$

ℓ_2 RIP:

$$(1 - \delta)\|x\|_{\ell_2}^2 \leq \|\mathbf{R}x\|_{\ell_2}^2 \leq (1 + \delta)\|x\|_{\ell_2}^2$$

$\mathbf{R} \sim$ Gaussian or subgaussian

ℓ_p RIP:

$$(1 - \delta)\|x\|_{\ell_1}^p \leq \|\mathbf{R}x\|_{\ell_p}^p \leq (1 + \delta)\|x\|_{\ell_1}^p$$

$\mathbf{R} \sim \alpha$ -stable

Why?: ℓ_1 norm is more robust and sparse inducing!

Why?

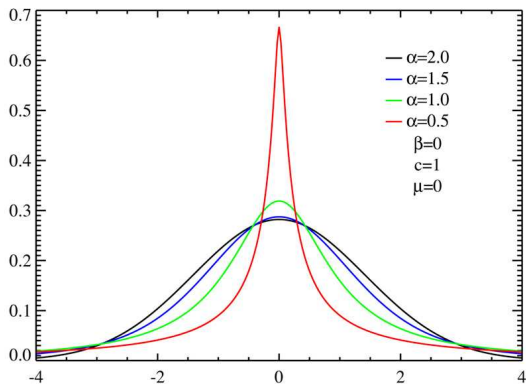
- α -stable distributions preserve the l_α norm. This one is more robust to noise, missing data, and outliers, than the l_2 norm.
- l_p minimization may be used using α -stable random projections. Furthermore, l_p minimization requires fewer measurements as p tends to zero.
- α -stable random projections are more robust to Gaussian noise and impulsive noise[†] than random projections with finite variance.

[†] If the impulsive noise has parameter α' , and $\alpha < \alpha'$, this is true.

α -STABLE RANDOM PROJECTIONS

The characteristic function for symmetric α -stable distributions ($S_{\alpha S}$), with location parameter zero and dispersion γ

$$\varphi(t) = e^{-\gamma|t|^\alpha}$$



Let $r_i \sim S\alpha S(0, \alpha, \gamma_i)$ and y be the linear combination of n i.i.d. $S\alpha S$ random variables: $y = r_1 x_1 + r_2 x_2 + \dots + r_n x_n$.
Stable law dictates that

$$y \sim S\alpha S(0, \alpha, \gamma)$$

$$\gamma = (|x_1 \gamma_1|^\alpha + |x_2 \gamma_2|^\alpha + \dots + |x_n \gamma_n|^\alpha)^{1/\alpha}.$$

When the $r_i \sim S\alpha S(0, \alpha, 1)$, the dispersion of y is the l_α norm of the scalars of the linear combination.

$$y \sim S\alpha S(0, \alpha, \gamma)$$

$$\gamma = \left(\sum_{i=1}^n |x_i|^\alpha \right)^{\frac{1}{\alpha}} = \|x\|_\alpha$$

Thus, α -stable distributions preserve the l_α norm.

α -STABLE RANDOM PROJECTIONS. CAUCHY RANDOM PROJECTIONS ($\alpha = 1$)

Cauchy random projections arise as a solution to ℓ_1 distance preservation in Dimensionality reduction applications.

Let $r_i \sim S\alpha S(0, 1, 1) \triangleq C(0, 1)$ then

$$y = \sum_{i=1}^n r_i x_i \sim C\left(0, \sum_{i=1}^n |x_i|\right)$$

For two vectors: $y_1 = \sum_{i=1}^n r_{1,i} x_{1,i}$ and $y_2 = \sum_{i=1}^n r_{2,i} x_{2,i}$

$$u_j = y_1 - y_2 \sim C\left(0, \sum_{i=1}^n |x_{1,i} - x_{2,i}|\right) = C(0, \|x_1 - x_2\|)$$

In Cauchy random projections, ℓ_1 distance estimation reduces to estimate the scale parameter γ from k i.i.d. samples $u_j \sim C(0, \gamma)$

$S\alpha S$ do not have finite variance, but its dispersion may be estimated using the fractional moment of order p . If $p < \alpha$, this moment is finite and it is given by:

$$E(|x|^p) = \frac{2^{p+1}\Gamma(-p/\alpha)}{\alpha\sqrt{\pi}\Gamma(-p/2)} \Gamma\left(\frac{p+1}{2}\right) \gamma^p.$$

From the latter expression, the following non-linear estimator follows

$$E(\|X\|_p^p) = E\left(\frac{1}{k} \sum_{i=1}^k |x_i|^p\right) = k \frac{2^{p+1}\Gamma(-p/\alpha)}{\alpha\sqrt{\pi}\Gamma(-p/2)} \Gamma\left(\frac{p+1}{2}\right) \gamma^p$$

The l_p^p norm of a vector –which components are α -stable realizations– is an estimate of the dispersion of its entries.

Theorem. Let \mathbf{R} be an $K \times N$ matrix whose elements are *i.i.d.* alpha-stable realizations with $1 \leq \alpha \leq 2$, which is multiplied by a scalar which depends only on K, α and ρ . Let $\|x\|_0 = s$, if $K \geq C_1 s \log(N/s) + C_2 s + C_3$ and $0 < \rho < \alpha/2$, then, for a given δ

$$(1 - \delta) \|x\|_\alpha^\rho < \|\mathbf{R}x\|_\rho^\rho < (1 + \delta) \|x\|_\alpha^\rho$$

holds with probability greater than $1 - 1/\binom{N}{s}$.

The proof of the theorem uses probabilistic methods. Rather than focusing on the values of the RIC, we choose a δ such that the RIP holds with high probability.

The proof will be divided in three lemmas and a proof that gathers the results obtained in the lemmas:

- **Lemma 1.** Probability that the RIP holds for a fixed x such that $x \in \mathbb{R}^s$ and $\|x\|_0 = s$.
- **Lemma 2.** Probability that the RIP fails for any $x \in \mathbb{R}^s$ having the same $\|x\|_\alpha^p$ such that $\|x\|_0 = s$.
- **Lemma 3.** Probability that the RIP fails for any submatrix of the matrix **R**.

Lemma 1. Let T be a set of indices with $|T| = s$ ($1 \leq s \leq N$), denote by X_T the set of discrete signals in \mathbb{R}^s that are zero outside of T . Let Ψ be an $K \times s$ submatrix of Φ formed with the columns whose indices are in T . If $0 < \rho < \alpha/2$ and for a $x \in X_T$, then, for a given $\eta > 0$

$$(1 - \eta)KC_\mu(\alpha, \rho) \|x\|_\alpha^\rho < \|\Psi x\|_\rho^\rho < (1 + \eta)KC_\mu(\alpha, \rho) \|x\|_\alpha^\rho$$

holds with probability lower-bounded by $1 - 2e^{-\frac{K\eta^2}{2(1+\eta)^2}C(\alpha, \rho)}$.

RESTRICTED- ρ ISOMETRY PROPERTY (RIP_ρ)

Proof : Let $y = \Psi x$, where x is a vector that belongs to X_T . The resultant vector y has K components, each one given by

$$y_i = \sum_{j=1}^s \psi_{ij} x_j \sim S\alpha S(0, \|x\|_\alpha).$$

The l_p^p norm of y is equal to

$$\|y\|_p^p = \|\Psi x\|_p^p = \sum_{i=1}^K \left| \sum_{j=1}^k \psi_{ij} x_j \right|^p = \sum_{i=1}^K |y_i|^p.$$

This implies that in average $\|\Psi x\|_p^p$ tends to

$$\begin{aligned} E(\|\Psi x\|_p^p) &= K \frac{2^{p+1} \Gamma(-p/\alpha)}{\alpha \sqrt{\pi} \Gamma(-p/2)} \Gamma\left(\frac{p+1}{2}\right) \|x\|_\alpha^p \\ &= K C_\mu(\alpha, p) \|x\|_\alpha^p. \end{aligned}$$

for $0 < p < \alpha$.

If $0 < p < \alpha/2$, the variance of the estimator $\|\Psi_X\|_p^p$ is finite as well

$$\begin{aligned} \text{Var}(\|\Psi_X\|_p^p) &= \text{Var}\left(\sum_{i=1}^K |y_i|^p\right) \\ &= KC_\mu(\alpha, 2p)\|X\|_\alpha^{2p} - KC_\mu(\alpha, p)^2\|X\|_\alpha^{2p} \\ &= KC_{\sigma^2}(\alpha, p)\|X\|_\alpha^{2p}. \end{aligned}$$

The distribution of $\|\Psi_X\|_p^p$ may be approximated with a Gaussian, however, a much better approximation may be done using a distribution with the same support of $\|\Psi_X\|_p^p$: the Inverse Gaussian.

If \mathbf{X} follows an inverse Gaussian distribution, it can be shown that \mathbf{X} has the following Chernoff bound:

$$P(|\mathbf{X} - \mu| < \eta\mu) \geq 1 - 2e^{-\frac{\eta^2}{2(1+\eta)}\left(\frac{\mu^2}{\sigma^2}\right)}.$$

It follows that

$$P(|\|\Psi_{\mathbf{X}}\|_p^p - \mu| < \eta\mu) \gtrsim 1 - 2e^{-\frac{K\eta^2}{2(1+\eta)^2}C(\alpha,p)} = 1 - P_{\alpha,p,K}(\eta).$$

Where $C(\alpha, p) = C_{\mu}(\alpha, p)^2 / C_{\sigma^2}(\alpha, p)$. If η is chosen properly, the RIP holds with high probability for a fixed \mathbf{x} .

Lemma 2. Let $0 < p < \alpha/2$, Ψ a $K \times s$ submatrix of Φ as in Lemma 1. Let $\delta > 0$ and choose $\epsilon > 0$ for a given η such that $\frac{\eta + \epsilon^p}{1 - \epsilon^p} \leq \delta$. Then for any $x \in \mathbb{R}^s$

$$(1 - \delta)KC_\mu(\alpha, p) \|x\|_\alpha^p < \|\Psi x\|_p^p < (1 + \delta)KC_\mu(\alpha, p) \|x\|_\alpha^p \quad (1)$$

holds with probability $\gtrsim 1 - (1 + 2/\epsilon)^s P_{\alpha, p, K}(\eta)$.

Lemma 3. Let Φ be an $K \times N$ matrix, $x \in \mathbb{R}^N$ and $\|x\|_0 = s$. Let $0 < p < \alpha/2$ and $\delta > 0$, then

$$(1 - \delta)KC_\mu(\alpha, p) \|x\|_\alpha^p < \|\Phi x\|_p^p < (1 + \delta)KC_\mu(\alpha, p) \|x\|_\alpha^p$$

with probability $\gtrsim 1 - (eN/s)^s(1 + 2/\epsilon)^s P_{\alpha,p,K}(\eta)$.

Proof of the Theorem. According to *Lemma 3*, we have

$$(1 - \delta)KC_\mu(\alpha, \rho) \|x\|_\alpha^p < \|\Phi x\|_p^p < (1 + \delta)KC_\mu(\alpha, \rho) \|x\|_\alpha^p$$

$$(1 - \delta) \|x\|_\alpha^p < \left\| (KC_\mu(\alpha, \rho))^{-1/\rho} \Phi x \right\|_p^p < (1 + \delta) \|x\|_\alpha^p.$$

Letting \mathbf{R} equal to $(KC_\mu(\alpha, \rho))^{-1/\rho} \Phi$, we obtain

$$(1 - \delta) \|x\|_\alpha^p < \|\mathbf{R}x\|_p^p < (1 + \delta) \|x\|_\alpha^p.$$

Then, the probability that the RIP fails is approximately less or equal than $(\frac{eN}{s})^s (1 + \frac{2}{\epsilon})^s P_{\alpha, \rho, K}(\eta)$.

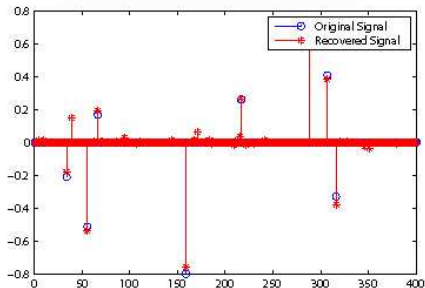
With a number of projections given by: $K \gtrsim C_1 s \log(N/s) + C_2 s + C_3$

COMPUTER SIMULATIONS

Sparse Signal Reconst. from Cauchy Random Proj. Noisy Case ($\alpha=1.2$)

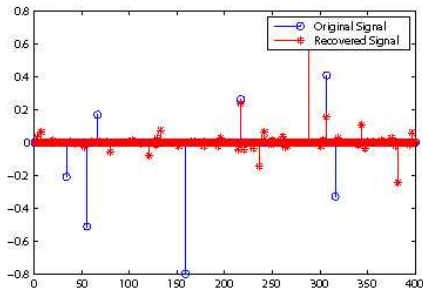
- Model: $y = Rx + \eta$; $f_{\eta}(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \varphi(t; 0, \alpha, \gamma) \exp^{-ixt} dt$.
- $n = 400$, $k = 100$, $s = 8$.
- $\ell_0 - LL$: ℓ_0 -regularized coordinate-descent Myriad-based regression.
- $\ell_1 - \ell_s$: ℓ_1 -regularized least squares.

$\ell_0 - LL$



$\gamma = 0.5$, MSE=-40.1dB

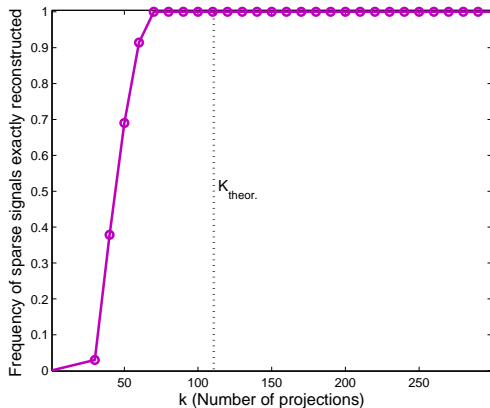
$\ell_1 - \ell_s$



$\gamma = 0.5$. MSE=-24.9dB

Number of required measurements. Noiseless case.

- Model: $y = Rx$.
- $n = 300$, $s = 6$.
- Reconstruction method: ℓ_0 -regularized coordinate-descent
Myriad-based regression.



- A new RIP for $S\alpha S$ random matrices with $1 \leq \alpha \leq 2$ has been formulated.
- Useful applications such as methods of robust statistics, reconstruction algorithms, among others.
- As $p \rightarrow 0$, less measurements are required for the RIP to hold with high probability.
- The restricted isometry constants (RIC) for this new RIP remain as an open problem.