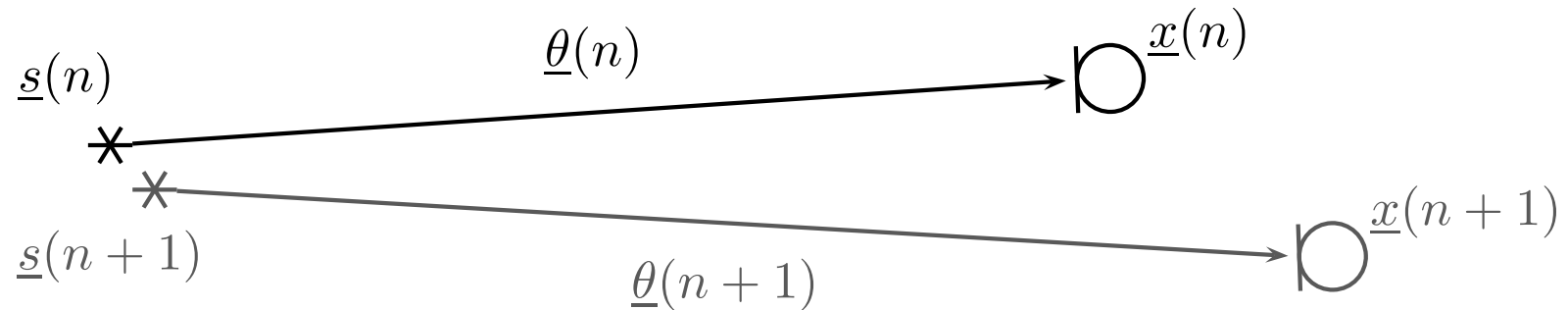

Recursive Estimation of Room Impulse Responses with Energy Conservation Constraints

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- **Goal:** Track time-varying room impulse response (RIR) between source and microphone



- **Signal model:**

$$\underline{x}(n) = \underline{\theta}(n) * \underline{s}(n) + \underline{z}(n), \quad n = 1, \dots, N$$

with

- $\underline{x}(n) \in \mathbb{R}^K$... measured microphone signal
- $\underline{\theta}(n) \in \mathbb{R}^M$... room impulse response
- $\underline{s}(n) \in \mathbb{R}^S$... known source signal
- $\underline{z}(n) \in \mathbb{R}^K$... Gaussian noise, $\underline{z}(n) \sim \mathcal{N}(\underline{0}, \mathbf{C}(n))$, temporally uncorr.

- **Signal model (in Matrix notation):**

$$\underline{x}(n) = \Theta(n)\underline{s}(n) + \underline{z}(n) = \mathbf{S}(n)\underline{\theta}(n) + \underline{z}(n), \quad n = 1, \dots, N$$

where $\Theta(n) \in \mathbb{R}^{K \times S}$ and $\mathbf{S}(n) \in \mathbb{R}^{K \times M}$ are Toeplitz matrices with

$$\Theta(n) = \begin{bmatrix} \underline{\theta}(n) & & & \mathbf{0} \\ & \underline{\theta}(n) & & \\ & & \ddots & \\ \mathbf{0} & & & \underline{\theta}(n) \end{bmatrix} \quad \mathbf{S}(n) = \begin{bmatrix} \underline{s}(n) & & & \mathbf{0} \\ & \underline{s}(n) & & \\ & & \ddots & \\ \mathbf{0} & & & \underline{s}(n) \end{bmatrix}$$

- **Popular solution approach:** Weighted least squares (WLS) estimator

$$\hat{\underline{\theta}}_{\text{WLS}}(n) = \arg \min_{\underline{\theta}} \sum_{i=0}^n \beta^{n-i} (\underline{x}(i) - \mathbf{S}(i)\underline{\theta})^T \mathbf{C}(i)^{-1} (\underline{x}(i) - \mathbf{S}(i)\underline{\theta})$$

where $0 \leq \beta \leq 1$ is the exponential forgetting factor

- **Idea:** Improve estimation performance by incorporating a priori information

- **Open Questions:**
 - **What a priori information (= constraint) can we exploit?**
 - ▶ Energy conservation constraint $\|\Theta(n)\underline{s}(n)\|^2 \leq \|\underline{s}(n)\|^2$

 - **How to efficiently include the constraint into tracking algorithm?**
 - ▶ Recursive constrained maximum likelihood (RCML) estimator
 - ▶ Recursive affine minimax (RAMX) estimator
 - ▶ Recursive minimum mean squared error (RMMSE) estimator



Energy Conservation Constraint

Room Impulse Response Estimators

Simulation Results

Summary

- Expressing the energy conservation mathematically, we need to ensure that

$$\|\Theta(n)\underline{s}(n)\|^2 = \underline{s}(n)^T \Theta(n)^T \Theta(n) \underline{s}(n) \leq \underline{s}(n)^T \underline{s}(n) = \|\underline{s}(n)\|^2$$

i.e.

“signal energy at microphone” \leq “signal energy at source”

which implies

$$\Theta(n)^T \Theta(n) - \mathbf{I} \preceq \mathbf{0}$$

- Condition has to hold for all signal lengths S
 - if condition fulfilled for particular S_0 then also fulfilled for all $S < S_0$
 - therefore, it is sufficient to consider case $S \rightarrow \infty$

- Two equivalent representations of the set $\Theta = \{\underline{\theta} : \Theta(n)^T \Theta(n) - \mathbf{I} \preceq \mathbf{0}\}$:

- **LMI representation:** Using Schur's lemma, the constraint is

$$\begin{bmatrix} \mathbf{I} & \Theta(n)^T \\ \Theta(n) & \mathbf{I} \end{bmatrix} \succeq \mathbf{0}$$

i.e. Θ is a convex set.

- **Frequency domain representation:**

- ▶ Use equivalence of the eigenvalues of $\Theta(n)^T \Theta(n)$ (bandlimited Toeplitz matrix) and corresponding circulant matrix for $S \rightarrow \infty$
- ▶ Constraint can be written as

$$|\Theta(\omega, n)|^2 = \left| \sum_{m=0}^{M-1} \theta_m(n) e^{-j\omega m} \right|^2 \leq 1 \quad \forall \omega \in [0, 2\pi)$$

where $\Theta(\omega, n)$ is the room frequency response

- All three recursive estimators are based on frequency domain representation
- Approximation of the frequency domain representation:

- Use DFT to evaluate $\Theta(\omega, n)$ at discrete frequencies ω_l

$$\omega_l = \frac{2\pi l}{L} \quad \text{with} \quad l = 0, \dots, L - 1$$

where $L \geq M$.

- Let $\mathbf{P} \in \mathbb{R}^{M \times L}$ be matrix consisting of the first M rows of \mathbf{I} . Then,

$$|\Theta(\omega_l, n)|^2 = \underline{\theta}(n)^T \mathbf{P} \underline{\tilde{f}}_l \underline{\tilde{f}}_l^H \mathbf{P}^T \underline{\theta}(n) = \underline{\tilde{\theta}}^T \underline{\tilde{f}}_l \underline{\tilde{f}}_l^H \underline{\tilde{\theta}} \leq 1 \quad \forall l = 0, \dots, \lfloor L/2 \rfloor$$

where $\underline{\tilde{f}}_l \in \mathbb{C}^L$ is the l th column of the DFT matrix $\mathbf{U} \in \mathbb{C}^{L \times L}$, i.e.

$$\underline{\tilde{f}}_l^H = \begin{bmatrix} 1 & e^{-j\omega_l} & \dots & e^{-j\omega_l(L-1)} \end{bmatrix}$$

- Consider the time-variant signal model

$$\underline{x}(n) = \mathbf{S}(n)\underline{\theta}_0(n) + \underline{z}(n), \quad n \geq 0$$

where it is assumed that $\underline{\theta}_0(n) \in \Theta \subset \mathbb{R}^M$

- Motivated by the WLS estimator, we reformulate this signal model into

$$\underline{x}_n = \mathbf{S}_n \underline{\theta}(n) + \underline{z}_n, \quad \underline{\theta}(n) \in \Theta$$

where \underline{x}_n , \mathbf{S}_n and \underline{z}_n are stacked versions of $\underline{x}(i)$, $\mathbf{S}(i)$ and $\underline{z}(i)$, i.e.

$$\begin{aligned} \underline{x}_n &= [\underline{x}(n)^T \quad \dots \quad \underline{x}(0)^T]^T \\ \mathbf{S}_n &= [\mathbf{S}(n)^T \quad \dots \quad \mathbf{S}(0)^T]^T, \\ \underline{z}_n &= [\underline{z}(n)^T \quad \beta^{-1/2}\underline{z}(n-1)^T \quad \dots \quad \beta^{-n/2}\underline{z}(0)^T]^T. \end{aligned}$$

- Noise $\underline{z}(n)$ Gaussian and temporally uncorrelated, i.e. $\underline{z}_n \sim \mathcal{N}(\underline{0}, \mathbf{C}_n)$ with

$$\mathbf{C}_n = \text{diag}(\mathbf{C}(n), \beta^{-1}\mathbf{C}(n-1), \dots, \beta^{-n}\mathbf{C}(0))$$

- How can we avoid estimators with growing computational complexity?
⇒ use concept of **sufficient statistics**
- A sufficient statistic for our signal model is given by

$$\underline{t}(\underline{x}_n) = (\mathbf{S}_n^T \mathbf{C}_n^{-1} \mathbf{S}_n)^{-1} \mathbf{S}_n^T \mathbf{C}_n^{-1} \underline{x}_n = \hat{\underline{\theta}}_{\text{WLS}}(n)$$

- Efficiently computed by **recursive weighted least squares algorithm** (RWLS)
 - in each time step, it updates

$$\begin{aligned} \hat{\underline{\theta}}_{\text{WLS}}(n) &= \underline{t}(\underline{x}_n) && \dots \quad \textit{sufficient statistic} \\ \mathbf{R}_n^{-1} &= (\mathbf{S}_n^T \mathbf{C}_n^{-1} \mathbf{S}_n)^{-1} && \dots \quad \textit{inverse correlaton matrix} \end{aligned}$$

- ⇒ **both quantities are needed by the three recursive estimators**
- ⇒ **estimators are based on RWLS and use constraint to improve performance**

- The RCML is given by

$$\begin{aligned}\hat{\underline{\theta}}_{\text{RCML}}(n) &= \arg \max_{\underline{\theta}} p(\underline{x}_n | \underline{\theta}) \quad \text{s.t.} \quad \underline{\theta} \in \Theta \\ &= \arg \min_{\underline{\theta}} (\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))^T \mathbf{R}_n (\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n)) \quad \text{s.t.} \quad \underline{\theta} \in \Theta\end{aligned}$$

At each time step, we have to check $\hat{\underline{\theta}}_{\text{WLS}}(n) \in \Theta$:

if yes: $\hat{\underline{\theta}}_{\text{RCML}}(n) = \hat{\underline{\theta}}_{\text{WLS}}(n)$

if no: find minimum of $(\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))^T \mathbf{R}_n (\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))$ on boundary of Θ

- For our RIR Tracking problem:

- Constraint is given by

$$|\Theta(\omega_l, n)|^2 = \underline{\theta}(n)^T \mathbf{P} \tilde{\underline{f}}_l \tilde{\underline{f}}_l^H \mathbf{P}^T \underline{\theta}(n) \leq 1 \quad \forall l = 0, \dots, \lfloor L/2 \rfloor$$

⇒ Quadratically constrained quadratic program (QCQP)

⇒ can be efficiently solved

- The RAMX is given by

$$\hat{\underline{\theta}}_{\text{RAMX}}(n) = (\mathbf{I} + \mathbf{M}(n)) \hat{\underline{\theta}}_{\text{WLS}}(n) + \underline{u}(n)$$

where $\mathbf{M}(n)$ and $\underline{u}(n)$ are the solution to the minimax problem

$$\min_{\mathbf{M}, \underline{u}} \max_{\underline{\theta} \in \Theta} \|\mathbf{M}\underline{\theta} + \underline{u}\|^2 + \text{tr} \{ (\mathbf{I} + \mathbf{M}) \mathbf{R}_n^{-1} (\mathbf{I} + \mathbf{M})^T \}$$

and the inverse correlation matrix $\mathbf{R}_n^{-1} = (\mathbf{S}_n^T \mathbf{C}_n^{-1} \mathbf{S}_n)^{-1}$ is updated by RWLS

- For our RIR Tracking problem:

- reduce computational complexity by using $\underline{u}(n) = \underline{0}$ and $\mathbf{M}(n) = \alpha \mathbf{I}$
- use S-procedure on Θ to reformulate optimization problem into SDP

- RMMSE recasts frequentist problem into Bayesian one:

- use uniform prior on Θ
- can be motivated by “maximum entropy principle”

- Using Bayesian sufficient statistics for our signal model, RMMSE is given by

$$\hat{\underline{\theta}}_{\text{RMMSE}}(n) = \frac{\int_{\Theta} \underline{\theta} \exp\left\{-\frac{1}{2}(\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))^T \mathbf{R}_n^{-1}(\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))\right\} d\underline{\theta}}{\int_{\Theta} \exp\left\{-\frac{1}{2}(\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))^T \mathbf{R}_n^{-1}(\underline{\theta} - \hat{\underline{\theta}}_{\text{WLS}}(n))\right\} d\underline{\theta}}$$

- Integrals are solved using Monte Carlo integration:

- Sampling from a posterior density using rejection sampling
- Posterior density is truncated Gaussian density (support given by Θ)

- For our RIR Tracking problem:

- use rejection sampling with $|\Theta(\omega_l, n)|^2 \leq 1$

- **Task:** Estimate RIR from moving source to fixed microphone

- **Setup:**
 - **Source**
 - ▶ Moves along a straight line, 10cm increment between two pos.
 - ▶ Overall 11 positions, i.e. source moves one meter in total
 - ▶ Source signal: Gaussian noise of length $S = 100$ samples

 - **Room:**
 - ▶ Room size: $3 \times 3 \times 2.3\text{m}$, $T_{60} = 120\text{ms}$
 - ▶ Image source model ($f_s = 12\text{kHz}$) results in RIRs with $M = 1241$ taps

 - **Estimators:**
 - ▶ For all three estimators: DFT length is $L = 2^{14} = 16\,384$
 - ▶ $I = 3\,000$ samples for RMMSE (MC approx. of integrals)

■ Definitions:

- Signal-to-noise ratio: $\text{SNR} = 10 \log_{10} \frac{\|\Theta(n)\underline{s}(n)\|^2}{(M+S-1)\sigma^2}$
- normalized error measure: $E = \frac{1}{N} \sum_{n=1}^N \frac{\|\hat{\underline{\theta}}(n) - \underline{\theta}(n)\|^2}{\|\underline{\theta}(n)\|^2}$

■ Instantaneous estimators:

SNR	WLS	CML	AMX	MMSE
5 dB	1.23×10^0	1.20×10^0	1.21×10^0	1.02×10^0
10 dB	3.88×10^{-1}	3.88×10^{-1}	3.87×10^{-1}	3.72×10^{-1}

■ Recursive estimators: $\beta_{\text{opt}} \approx 0.55$ for SNR = 5dB; $\beta_{\text{opt}} \approx 0.28$ for SNR = 10dB

SNR	RWLS	RCML	RAMX	RMMSE
5 dB	7.02×10^{-1}	6.97×10^{-1}	6.99×10^{-1}	6.52×10^{-1}
10 dB	2.98×10^{-1}	2.98×10^{-1}	2.98×10^{-1}	2.94×10^{-1}

- Recursive RIR estimation with energy conservation constraint was studied
 - Constraint helps to improve performance
 - RMMSE estimator with uniform prior performs best

- To further improve performance: incorporate additional a priori information

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Thank you for your attention

Backup

- The condition $\Theta(n)^T \Theta(n) - \mathbf{I} \preceq \mathbf{0}$ requires all eigenvalues of $\Theta(n)^T \Theta(n)$ to be smaller or equal to 1

- Let

$$r(n) = \begin{cases} \sum_{m=0}^{M-1-|n|} \theta_m(n) \theta_{|n|+m}(n) & |n| \leq M-1 \\ 0 & \text{otherwise} \end{cases}$$

denote the (unnormalized) auto-correlation function of the RIR $\underline{\theta}(n)$

- Then, the matrix $\Theta(n)^T \Theta(n) \in \mathbb{R}^{S \times S}$ can be written as a symmetric Toeplitz matrix where the first row is given by $r(n)$ for $n = 0, \dots, S-1$, i.e. $\Theta(n)^T \Theta(n)$ is bandlimited
- Using the asymptotic equivalence of the eigenvalues of a bandlimited Toeplitz matrix and the corresponding circulant matrix, it follows

$$|\Theta(\omega, n)|^2 = \left| \sum_{m=0}^{M-1} \theta_m(n) e^{-j\omega m} \right|^2 \leq 1 \quad \forall \omega \in [0, 2\pi)$$

- Epigraphic form of the AMX optimization problem:

$$\min_{\mathbf{M}, \underline{u}, \tau} \tau$$

subject to

$$\begin{bmatrix} \underline{\theta} \\ 1 \end{bmatrix}^T \begin{bmatrix} -\mathbf{M}^T \mathbf{M} & -\mathbf{M}^T \underline{u} \\ -\underline{u}^T \mathbf{M} & \tau - \text{tr} \{ (\mathbf{I} + \mathbf{M}) \mathbf{R}_n^{-1} (\mathbf{I} + \mathbf{M})^T \} - \underline{u}^T \underline{u} \end{bmatrix} \begin{bmatrix} \underline{\theta} \\ 1 \end{bmatrix} \geq 0 \quad \forall \underline{\theta} \in \Theta$$

- **S-procedure:** sufficient condition for the statement

$$\text{for all } \underline{z}: \underline{z}^T \mathbf{F}_0 \underline{z} \geq 0, \dots, \underline{z}^T \mathbf{F}_{\tilde{L}} \underline{z} \geq 0 \Rightarrow \underline{z}^T \mathbf{G} \underline{z} \geq 0$$

to be true is the existence of $\lambda_0, \dots, \lambda_{\tilde{L}} \geq 0$ such that $\mathbf{G} \succeq \lambda_0 \mathbf{F}_0 + \dots + \lambda_{\tilde{L}} \mathbf{F}_{\tilde{L}}$.

■ Final optimization problem:

$$\min_{\substack{\lambda_0 \geq 0, \dots, \lambda_{\tilde{L}} \geq 0 \\ \tau, \alpha, x}} \tau$$

subject to

$$\lambda_l \geq x/L \quad (l = 0, \dots, \tilde{L}),$$

$$\begin{bmatrix} x & \alpha \\ \alpha & 1 \end{bmatrix} \succeq \mathbf{0},$$

$$\tau - (1 + 2\alpha + x) \operatorname{tr}\{\mathbf{R}_n^{-1}\} \geq \lambda_0 + \dots + \lambda_{\tilde{L}}.$$

■ Definition: (introduced by Kolmogorov)

- $\underline{t}(\underline{x})$ is **Bayesian sufficient statistic** if for all priors $p(\underline{\theta})$ the a posteriori density $p(\underline{\theta}|\underline{x})$ can be written as $p(\underline{\theta}|\underline{x}) = p(\underline{\theta}|\underline{t}(\underline{x}))$.

⇒ Given $\underline{t}(\underline{x})$, no additional information can be inferred from \underline{x}

■ Concept directly related to classical sufficiency:

- using Fisher-Neyman factorization theorem: Classical sufficiency implies Bayesian sufficiency

$$p(\underline{\theta}|\underline{x}) = \frac{p(\underline{\theta})p(\underline{x}|\underline{\theta})}{\int p(\underline{\theta})p(\underline{x}|\underline{\theta})d\theta} = \frac{p(\underline{\theta})g(\underline{t}(\underline{x}), \underline{\theta})}{\int p(\underline{\theta})g(\underline{t}(\underline{x}), \underline{\theta})d\theta} = p(\underline{\theta}|\underline{t}(\underline{x}))$$

- converse is also true under some regularity conditions